gamboostLSS: boosting generalized additive models for location, scale and shape

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Motivation: Munich rental guide

Aim:

• Provide precise point predictions and prediction intervals for the net-rent of flats in the city of Munich.

Data:

- Covariates: 325 (mostly) categorical, 2 continuous and 1 spatial
- Observations: 3016 flats

Problem:

• Heteroscedasticity found in the data

Idea

Model not only the expected mean but also the variance \Rightarrow GAMLSS

The GAMLSS model class

Generalized Additive Models for Location, Scale and Shape

$$g_1(\mu) = \eta_\mu = \beta_{0\mu} + \sum_{j=1}^{p_1} f_{j\mu}(x_j) \qquad \text{``location''}$$
$$g_2(\sigma) = \eta_\sigma = \beta_{0\sigma} + \sum_{j=1}^{p_2} f_{j\sigma}(x_j) \qquad \text{``scale''}$$
$$\vdots$$

- Introduced by Rigby and Stasinopoulos (2005)
- Flexible alternative to generalized additive models (GAM)
- Up to four distribution parameters are regressed on the covariates.
- Every distribution parameter is modeled by its own predictor and an associated link function $g_k(\cdot)$.

Current fitting algorithm

The gamlss package

Fitting algorithms for a large amount of distribution families are provided by the **R** package gamlss (Stasinopoulus and Rigby, 2007).

- Estimation is based on a penalized likelihood approach.
- Modified versions of back-fitting (as for conventional GAMs) are used.

These algorithms work remarkably well in many applications, but:

- It is not feasible for high-dimensional data $(p \gg n)$.
- No spatial effects are implemented.
- Variable selection is based on generalized AIC, which is known to be unstable.
 - ▷ "More work needs to be done here" (Stasinopoulus and Rigby, 2007).

Optimization problem for GAMLSS

- The task is to model the distribution parameters of the conditional density $f_{\rm dens}(y|\mu,\sigma,\nu,\tau)$
 - > The optimization problem can be formulated as

$$\underbrace{(\hat{\mu},\hat{\sigma},\hat{\nu},\hat{\tau})}_{\boldsymbol{\theta}} \longleftarrow \operatorname{argmin}_{\eta_{\mu},\eta_{\sigma},\eta_{\nu},\eta_{\tau}} \mathbb{E}_{Y,X} \left[\rho \big(Y, \underbrace{\eta_{\mu}(X),\eta_{\sigma}(X),\eta_{\nu}(X),\eta_{\tau}(X)}_{\boldsymbol{\eta}} \big) \right]$$

with loss function $\rho=-l$, i.e., the negative log-likelihood of the response distribution:

$$l = \sum_{i=1}^{n} \log \left[f_{\mathsf{dens}}(y_i | \boldsymbol{\theta}_i) \right] = \sum_{i=1}^{n} \log \left[f_{\mathsf{dens}}(y_i | \mu_i, \sigma_i, \nu_i, \tau_i) \right]$$

Maximum likelihood approach

Alternative to ML: ▷ Component-wise boosting Boosting

- minimizes empirical risk (e.g., negative log likelihood)
- in an iterative fashion
- via functional gradient descent (FGD).

In boosting iteration m+1

• Compute (negative) gradient of the loss function and plug in the current estimate

$$u_i^{[m+1]} = - \left. \frac{\partial \rho(y_i, \eta)}{\partial \eta} \right|_{\eta = \hat{\eta}_i^{[m]}}$$

• Estimate $u_i^{[m+1]}$ via base-learners (i.e., simple regression models)

• Update: use only the **best-fitting base-learner**; add a small fraction ν of this estimated base-learner (e.g., 10%) to the model

▷ Variable selection intrinsically within the fitting process

- Boosting was recently extended to risk functions with multiple components (Schmid et al., 2010)
- Idea ▷ Use partial derivatives instead of gradient
- Specify a set of base-learners one base-learner per covariate
- Fit each of the base-learners separately to the partial derivatives
- Cycle through the partial derivatives within each boosting step

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$$\frac{\partial \rho}{\partial \eta_{\mu}}(y_i, \hat{\mu}^{[m]}, \hat{\sigma}^{[m]}, \hat{\nu}^{[m]}, \hat{\tau}^{[m]}) \qquad \xrightarrow[\text{best fitting BL}}^{\text{update}} \hat{\eta}_{\mu}^{[\mathbf{m+1}]} \Longrightarrow \hat{\mu}^{[\mathbf{m+1}]}$$

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$$\begin{array}{ll} \frac{\partial\rho}{\partial\eta_{\mu}}(y_{i},\hat{\mu}^{[m]},\hat{\sigma}^{[m]},\hat{\nu}^{[m]},\hat{\tau}^{[m]}) & \stackrel{\text{update}}{\underset{\text{best fitting BL}}{\longrightarrow}} \hat{\eta}_{\mu}^{[\mathbf{m+1}]} \Longrightarrow \hat{\mu}^{[\mathbf{m+1}]} , \\ \frac{\partial\rho}{\partial\eta_{\sigma}}(y_{i},\hat{\mu}^{[\mathbf{m+1}]},\hat{\sigma}^{[m]},\hat{\nu}^{[m]},\hat{\tau}^{[m]}) & \stackrel{\text{update}}{\underset{\text{best fitting BL}}{\longrightarrow}} \hat{\eta}_{\sigma}^{[\mathbf{m+1}]} \Longrightarrow \hat{\sigma}^{[\mathbf{m+1}]} , \end{array}$$

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Variable selection and shrinkage

- The main tuning parameter are the stopping iterations $m_{\text{stop},k}$. They control variable selection and the amount of shrinkage.
 - If boosting is stopped before convergence only the most important variables are included in the final model.
 - Variables that have never been selected in the updated step, are excluded.
 - Due to the small increments added in the update step, boosting incorporates shrinkage of effect sizes (compare to LASSO), leading to more stable predictions.
- For large $m_{\text{stop},k}$ boosting converges to the same solution as the original algorithm (in low-dimensional settings).
- The selection of $m_{\text{stop},k}$ is normally based on resampling methods, optimizing the predictive risk.

Data example: Munich rental guide

To deal with heteroscedasticity, we chose a three-parametric t-distribution with

$$\mathbb{E}(y) = \mu$$
 and $\mathbb{V}ar(y) = \sigma^2 \frac{\mathrm{d} \mathbf{f}}{\mathrm{d} \mathbf{f} - 2}$

For each of the parameters $\mu,\,\sigma,$ and df, we consider the candidate predictors

$$\begin{split} \eta_{\mu_i} &= \beta_{0\mu} + x_i^\top \beta_\mu + f_{1,\mu}(\mathsf{size}_i) + f_{2,\mu}(\mathsf{year}_i) + f_{\mathsf{spat},\mu}(s_i) \ , \\ \eta_{\sigma_i} &= \beta_{0\sigma} + x_i^\top \beta_\sigma + f_{1,\sigma}(\mathsf{size}_i) + f_{2,\sigma}(\mathsf{year}_i) + f_{\mathsf{spat},\sigma}(s_i) \ , \\ \eta_{\mathsf{df}_i} &= \beta_{\mathsf{0df}} + x_i^\top \beta_{\mathsf{df}} + f_{1,\mathsf{df}}(\mathsf{size}_i) + f_{2,\mathsf{df}}(\mathsf{year}_i) + f_{\mathsf{spat},\mathsf{df}}(s_i) \ . \end{split}$$

Base-learners

- Categorical variables: Simple linear models
- Continuous variables: P-splines
- Spatial variable: Gaussian MRF (Markov random fields)

Package gamboostLSS

- Boosting for GAMLSS models is implemented in the R package gamboostLSS (▷ now available on CRAN).
- Package relies on the well tested and mature boosting package **mboost**.
- Lots of the **mboost** infrastructure is available in **gamboostLSS** as well (e.g., base-learners & convenience functions).

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Now let's start and have a short look at some code!

Package gamboostLSS

- > library("gamboostLSS")

(Simplified) code to fit the model

```
> ## Load data first, and load boundary file for spatial effects
> ## Now set up formula:
> form <- paste(names(data)[1], " ~ ",</pre>
               paste(names(data)[-c(1, 327, 328, 329)], collapse = " + "),
               " + bbs(wfl) + bbs(bamet) + bmrf(region, bnd = bound)")
> form <- as.formula(form)</pre>
> form
nmqms ~ erstbezg + dienstwg + gebmeist + gebgruen + hzkohojn +
   .... +
   bbs(wfl) + bbs(bamet) + bmrf(region, bnd = bound)
> ## Fit the model with (initially) 100 boosting steps
> mod <- gamboostLSS(formula = form, families = StudentTLSS(),</pre>
                    control = boost_control(mstop = 100,
                                           trace = TRUE).
                    baselearner = bols,
                    data = data)
[ 1] ..... -- risk: 3294.323
[ 41] ..... -- risk: 3091.206
[ 81] .....
Final risk: 3038,919
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                               gamboostLSS
                                                                  12 / 18
```

(Simplified) code to fit the model (ctd.)

```
> ## optimal number of boosting iterations fund by 3-dimensional
> ## cross-validation on a logarithmic grid resulted in
> ## 750 (mu), 108 (sigma), 235 (df) steps;
> ## Let model run until these values:
> mod[c(750, 108, 235)]
>
> ## Let's look at the number of variables per parameter:
> sel <- selected(mod)</pre>
> lapply(sel, function(x) length(unique(x)))
$m11
[1] 115
$sigma
[1] 31
$df
[1] 7
```

> ## (Very) sparse model (only 115, 31 and 5 base-learners out of 328)

(Simplified) code to fit the model (ctd.)

- > ## Now we can look at the estimated parameters
- > ## e.g., the effect of roof terrace on the mean

```
> coef(mod, which = "dterasn", parameter = "mu")
```

- \$'bols(dterasn)'
- (Intercept) dterasn
- -0.004254606 0.293792997
- + xlab = "flat size (in square meters)", type = "l")



flat size (in square meters)

Results

Results: spatial effects



Estimated spatial effects obtained for the high-dimensional GAMLSS for distribution parameters μ and σ . For the third parameter df, the corresponding variable was not selected.

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gamboostLSS

Results: prediction intervals



95% prediction intervals based on the quantiles of the modeled conditional distribution. Coverage probability GAMLSS 93.93% (92.07-95.80); coverage probability GAM 92.23% (89.45-94.32).

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Summary

- As gamboostLSS relies on mboost, we have a well tested, mature back end.
- The base-learners offer great flexibility when it comes to the **type of effects** (linear, non-linear, spatial, random, monotonic, ...).
- Boosting is feasible even if $p \gg n$.
- Variable selection is included in the fitting process. Additional shrinkage leads to more stable results.

The algorithm is implemented in the **R** add-on package **gamboostLSS now** available on CRAN.

Further literature

- Mayr, A., N. Fenske, B. Hofner, T. Kneib and M. Schmid, (2010). GAMLSS for high-dimensional data – a flexible approach based on boosting. *Journal* of the Royal Statistical Society, Series C (Applied Statistics), accepted.
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 - Bühlmann, P. and Hothorn, T., (2007). Boosting algorithms: Regularization, prediction and model fitting. *Statistical Science*, **22**, 477-522.
 - Rigby, R. A. and Stasinopoulos, D. M., (2005). Generalized additive models for location, scale and shape. *Applied Statistics*, **54**, 507-554.
 - Schmid, M., S. Potapov, A. Pfahlberg and T. Hothorn, (2010). Estimation and regularization techniques for regression models with multidimensional prediction functions . *Statistics and Computing*, **20**, 139-150.